

**Zeitschrift:** Elemente der Mathematik  
**Herausgeber:** Schweizerische Mathematische Gesellschaft  
**Band:** 39 (1984)  
**Heft:** 2

**Artikel:** A remark on buffon's needle problem  
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**DOI:** <https://doi.org/10.5169/seals-38015>

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## A Remark on Buffon's Needle Problem

Let  $G_n$  be a grid in the Euclidean space  $\mathbb{R}^n$  with coordinates  $x_1, x_2, \dots, x_n$  determined by hyperplanes parallel to the hyperplanes with the equation  $x_i = 0$  separated by a distance of  $2L$ .

In his generalization of Buffon's Needle Problem to  $n$  dimensions Stoka [2] gives an expression for the probability that a segment  $\omega$  of length  $L$  which will be "thrown" in a random fashion into the  $\mathbb{R}^n$  cuts the grid  $G_n$ .

Let  $A_j$  be the event: the segment  $\omega$  cuts a hyperplane of  $G_n$  parallel to the hyperplane with the equation  $x_j = 0$ . In [2] and [3] Stoka shows that the probability for the event  $A_j$  is

$$P(A_j) = \frac{\Gamma\left(\frac{n}{2}\right)}{(n-1) \Gamma\left(\frac{n-1}{2}\right) \sqrt{\pi}} \quad (j = 1, 2, \dots, n) \tag{1}$$

(formula (9) in [3] is a misprint) and finds for the estimator

$$\hat{P}_n = \frac{1}{nN} \sum_{j=1}^n (\text{number of times } A_j \text{ occurs in } N \text{ independent trials})$$

of  $P(A_j)$  the variance

$$D^2(\hat{P}_n) = \frac{1}{nN} \theta_n \tag{2}$$

with

$$\begin{aligned} \theta_n = & \frac{(2n-1) \Gamma\left(\frac{n}{2}\right)}{(n-1) \Gamma\left(\frac{n-1}{2}\right) \sqrt{\pi}} - \frac{n \left[ \Gamma\left(\frac{n}{2}\right) \right]^2}{(n-1)^2 \left[ \Gamma\left(\frac{n-1}{2}\right) \right]^2 \pi} - n + 1 \\ & + \frac{2^{n-2} (n-1) \Gamma\left(\frac{n}{2}\right)}{\pi^{3/2} (n-2)!} \left[ \left( \pi + \frac{1}{2n} \right) \Gamma\left(\frac{n-1}{2}\right) - \frac{2\sqrt{\pi}}{n-1} \Gamma\left(\frac{n}{2}\right) \right] \end{aligned}$$

([3], formula (15)). On the other hand he calculates for the estimator  $\hat{P}_1 = M^{-1}$  (number of times  $A_1$  occurs in  $M$  independent trials) of  $P(A_j)$  the variance

$$D^2(\hat{P}_1) = \frac{1}{M} \frac{\Gamma\left(\frac{n}{2}\right)}{(n-1) \Gamma\left(\frac{n-1}{2}\right) \sqrt{\pi}} \left[ 1 - \frac{\Gamma\left(\frac{n}{2}\right)}{(n-1) \Gamma\left(\frac{n-1}{2}\right) \sqrt{\pi}} \right] \tag{3}$$

([3], formula (16)). For the case  $n = 2$  compare Schuster [1]. If we set both variances (2) and (3) equal we receive

$$M = \sigma(n) n N \tag{4}$$

with

$$\sigma(n) = \frac{\Gamma\left(\frac{n}{2}\right) \left[ (n-1) \Gamma\left(\frac{n-1}{2}\right) \sqrt{\pi} - \Gamma\left(\frac{n}{2}\right) \right]}{(n-1)^2 \left[ \Gamma\left(\frac{n-1}{2}\right) \right]^2 \pi \theta_n} . \tag{5}$$

Thus,  $N$  independent trials in  $G_n$  give the same information about the probability (1) as  $\sigma(n) n N$  independent trials in a grid determined by hyperplanes parallel to the hyperplane with the equation  $x_1 = 0$ . Schuster [1] finds  $\sigma(2) \approx 1.1114$ . Stoka [3] calculates  $\sigma(3) \approx 1.1121$  and  $\sigma(4) \approx 1.1039$  and conjectures that  $\lim_{n \rightarrow \infty} \sigma(n) = 1$ . In this note we give an asymptotic expression for  $\sigma(n)$  and show that Stoka's conjecture is true.

**Theorem.** *It holds*

$$\sigma(n) = \frac{\pi n! - 2^{n-2} n \left[ \Gamma\left(\frac{n}{2}\right) \right]^2}{\pi n! - 2^{n-2} n^2 \left[ \Gamma\left(\frac{n}{2}\right) \right]^2 + 2^{n-3} (n-1)^3 \left[ \Gamma\left(\frac{n-1}{2}\right) \right]^2} , \tag{6}$$

$$\sigma(n) \cong \frac{\sqrt{2\pi} - n^{-1/2} \left(\frac{n-2}{n}\right)^{n-1} e^2}{\sqrt{2\pi} - n^{1/2} \left(\frac{n-2}{n}\right)^{n-1} e^2 + n^{1/2} \left(\frac{n-1}{n}\right)^3 \left(\frac{n-3}{n}\right)^{n-2} e^3} \quad (n \rightarrow \infty) \tag{7}$$

and

$$\lim_{n \rightarrow \infty} \sigma(n) = 1 . \tag{8}$$

**Proof:** First we simplify the expression (5). For abbreviation we set

$a = a_n = \Gamma\left(\frac{n-1}{2}\right)$  and  $b = b_n = \Gamma\left(\frac{n}{2}\right)$ . Then by (5)

$$\sigma(n) = \frac{(n-1) \sqrt{\pi} a b - b^2}{(n-1)^2 \pi \left\{ \frac{(2n-1) a b}{(n-1) \sqrt{\pi}} - \frac{n b^2}{(n-1)^2 \pi} - (n-1) a^2 + \right.}$$

$$\left. + \frac{2^{n-2} (n-1) \left(\pi + \frac{1}{2n}\right) a^3 b}{\pi^{3/2} (n-2)!} - \frac{2^{n-1} a^2 b^2}{\pi (n-2)!} \right\} .$$

The formula

$$\Gamma(x) \Gamma\left(x + \frac{1}{2}\right) = \frac{\sqrt{\pi}}{2^{2x-1}} \Gamma(2x)$$

gives us

$$ab = \frac{\sqrt{\pi}}{2^{n-2}} (n-2)!$$

and therefore

$$\begin{aligned} \sigma(n) &= \frac{2^{-(n-2)} \pi (n-1)! - b^2}{2^{-(n-2)} \pi (n-1)! (2n-1) - nb^2 + \frac{(n-1)^3}{2n} a^2 - 2^{-(n-3)} \pi (n-1)! (n-1)} \\ &= \frac{\pi n! - 2^{n-2} n b^2}{\pi n! - 2^{n-2} n^2 b^2 + 2^{n-3} (n-1)^3 a^2} \end{aligned}$$

which is (6).

By Stirling's formula for the Gamma-function it holds

$$\log \Gamma(x+1) = \left(x + \frac{1}{2}\right) \log x - x + c + o\left(\frac{1}{x}\right) \quad (x \rightarrow +\infty)$$

with

$$c = \frac{1}{2} \log(2\pi).$$

This implies

$$\begin{aligned} \sigma(n) &\cong \frac{\pi n^{n+1/2} e^{-n+c} - 2^{n-2} n \left(\frac{n-2}{2}\right)^{n-1} e^{-(n-2)+2c}}{\pi n^{n+1/2} e^{-n+c} - 2^{n-2} n^2 \left(\frac{n-2}{2}\right)^{n-1} e^{-(n-2)+2c} + 2^{n-3} (n-1)^3 \left(\frac{n-3}{2}\right)^{n-2} e^{-(n-3)+2c}} \\ &= \frac{\sqrt{2\pi} - n^{-1/2} \left(\frac{n-2}{n}\right)^{n-1} e^2}{\sqrt{2\pi} - n^{1/2} \left(\frac{n-2}{n}\right)^{n-1} e^2 + n^{1/2} \left(\frac{n-1}{n}\right)^3 \left(\frac{n-3}{n}\right)^{n-2} e^3} \end{aligned}$$

which is (7).

Evidently the numerator of (7) converges to  $\sqrt{2\pi}$  for  $n \rightarrow \infty$ . It is sufficient to show that

$$d = \lim_{n \rightarrow \infty} n^{1/2} \left[ \left( \frac{n-2}{n} \right)^{n-1} - \left( \frac{n-1}{n} \right)^3 \left( \frac{n-3}{n} \right)^{n-2} e \right] = 0.$$

Now

$$|d| \leq \lim_{n \rightarrow \infty} n^{1/2} \left( \frac{n-1}{n} \right)^3 \left( \frac{n-3}{n} \right)^{n-2} \cdot \left\{ \left| \left( \frac{n}{n-1} \right)^3 \left( \frac{n-2}{n} \right) - 1 \right| \left( \frac{n-2}{n-3} \right)^{n-2} + \left| \left( \frac{n-2}{n-3} \right)^{n-2} - e \right| \right\}.$$

Because of

$$\left( \frac{n}{n-1} \right)^3 \frac{n-2}{n} - 1 = 0 \left( \frac{1}{n} \right) \quad (n \rightarrow \infty)$$

it remains to show that

$$\left( 1 + \frac{1}{n} \right)^{n+1} = e + 0 \left( \frac{1}{n} \right) \quad (n \rightarrow \infty).$$

To see this we set  $x_i = i/n$  ( $i = 0, 1, \dots, n$ ) and Taylor's formula gives us

$$e^{x_i} = \left( 1 + \frac{1}{n} \right) e^{x_{i-1}} + \frac{1}{2n^2} e^{\xi_i} \quad \text{with} \quad x_{i-1} < \xi_i < x_i.$$

This implies

$$\begin{aligned} e = e^{x_n} &= \left( 1 + \frac{1}{n} \right)^n e^{x_0} + \frac{1}{2n^2} \sum_{i=1}^n \left( 1 + \frac{1}{n} \right)^{n-i} e^{\xi_i} \\ &= \left( 1 + \frac{1}{n} \right)^n + 0 \left( \frac{1}{n} \right) \quad (n \rightarrow \infty). \end{aligned}$$

Thus the Theorem is proved.

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